

Integer Programming for Generalized Causal Bootstrap Designs

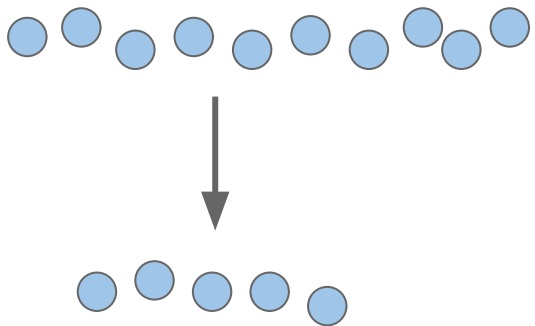
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Two sources of uncertainty in randomized experiments



Sampling Uncertainty: What if we had drawn a different set of units?



Design Uncertainty: What if we had assigned different units to treatment vs. control?

Claim: In randomized experiments, it is important to assess design uncertainty – and doing so may give tighter CI's than incorrectly applying the standard bootstrap.

Contributions

Previous work

A causal bootstrap procedure to estimate **design uncertainty** in the **difference in means estimator** under **complete randomization**.

This work

- **Integer programming** implementation of the causal bootstrap.
- Extensions to **linear-** and **quadratic-in-treatment estimators**.
- Extensions to **general randomized designs** where treatment probabilities and covariance are known.
- Application to **geographic experiments**.

Measuring design uncertainty

What we want

"If I were to repeat this experiment many times, what would the 5th and 95th percentiles on the test statistic be?"

i	$Y_i(0)$	$Y_i(1)$
1	9	10
2	10	11.5
3	8.5	10
4	12	13



i	$Y_i(0)$	$Y_i(1)$
1	?	10
2	?	11.5
3	8.5	?
4	12	?

$$\hat{\tau} = -0.5$$

i	$Y_i(0)$	$Y_i(1)$
1	?	10
2	10	?
3	?	10
4	12	?

$$\hat{\tau} = -1$$

i	$Y_i(0)$	$Y_i(1)$
1	9	?
2	?	11.5
3	?	10
4	12	?

$$\hat{\tau} = 0.25$$

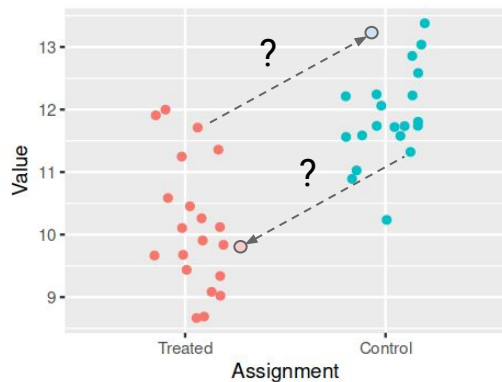
What we have

A single experimental observation

i	$Y_i(0)$	$Y_i(1)$
1	9	?
2	?	11.5
3	?	10
4	12	?

The Causal Bootstrap [Imbens-Menzel '21]

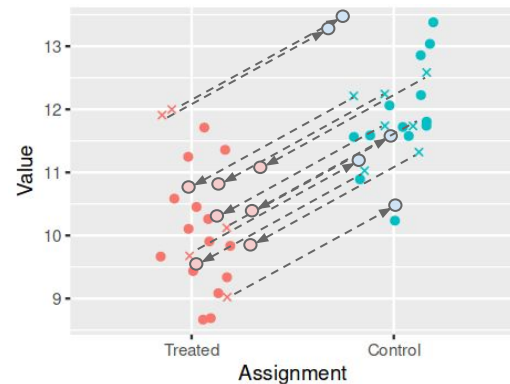
To generate one bootstrap replicate...



i	$Y_i(0)$	$Y_i(1)$
1	9	10
2	10	11.5
3	8.5	10
4	12	13



i	$Y_i(0)$	$Y_i(1)$
1	9	10
2	10	11.5
3	8.5	10
4	12	13



1) Impute the outcome of each unit under the unobserved condition

2) Draw a new randomization of the same units, using observed data if available and imputed data otherwise

Key Idea [Robins '88]: Impute via the joint distribution that maximizes the variance of your estimator, subject to matching the observed control and treatment marginal outcomes.

For diff-in-means + complete randomization, this is the assortative copula [Aronow et al '14].

Integer Programming Formulation

- Let $Z_i \in \{0, 1\}$ denote the binary treatment indicator for unit i .
- Let $X_{ik}^{(a)} \in \{0, 1\}$ be a binary optimization variable, indicating whether unit i has outcome k under treatment status a .
- \mathbf{Q} is a symmetric positive definite matrix that depends on $\Pr(Z_i = 1)$ and $\text{Cov}(Z_i, Z_j)$ for all pairs of units i, j .

$$\max \quad \mathbf{X}^T \mathbf{Q} \mathbf{X} \quad (4)$$

————→ maximize estimator variance

$$\text{s.t.} \quad \mathbf{X} \in \{0, 1\}^{N \cdot |\mathcal{Y}|}$$

$$\forall i, k, \quad X_{ik}^{(Z_i)} = 1 \text{ iff } Y_i^{\text{obs}} = y_k \quad (a)$$

————→ assign unit to observed outcome

$$\forall a, i, k, \quad X_{ik}^{(a)} = 0 \text{ iff } y_k \notin \text{supp}(F_a) \quad (b)$$

————→ avoid outcomes not in support

$$\forall a, i, \quad \sum_{k=1}^K X_{ik}^{(a)} = 1 \quad (c)$$

————→ each unit is assigned to at most one outcome for $a = 0, 1$

$$\forall a, k, b, \quad (-1)^b \sum_{i=1}^N X_{ik}^{(a)} \left(\frac{Z_i}{N_1} - \frac{1 - Z_i}{N_0} \right) \leq \epsilon \quad (d)$$

————→ marginals must stay the same

Simulation Result

We test our method on a **simulated geographical experiment** estimating the effect of an intervention on countries' GDP.

We consider two designs: **complete randomization** and **matched pairs**, and two estimators: **difference in means** and **model imputation**.

Our causal bootstrap achieves **almost nominal coverage**, with **up to 45% narrower confidence intervals** compared to the standard bootstrap.