



#### Background

- based on data collected from a potentially different and unknown policy.
- Among various algorithms for OPE, fitted Q-evaluation (FQE) is arguably one of the most popular algorithms.
- > FQE has demonstrated significant empirical success in many applications, the theoretical analysis of FQE is less explored in current literature.

## Set up Finite-horizon episodic Markov Decision Process (MDP): T is the length of horizon, S is the state space, A is the action space, $Pr_t(\cdot \mid s, a)$ representing the transition kernel (probability) at step t given the state $s \in S$ and the action $a \in A$ , $R_t$ is the immediate reward at step t. • Given pre-collected training data consist of *n* independent and identically distributed trajectories $\mathcal{D}_n = \left\{ \left\{ \left(S_{i,t}, A_{i,t}, R_{i,t} \right) \right\}_{1 \le t < T} \right\}_{1 \le i \le n}$ , OPE aims to estimate the value of $\pi$ defined as $u(\pi) = \mathbb{E}^{\pi} \left| \sum_{t=1}^{T} \mathsf{R}_t \right| \,.$ • Define $\rho_t^{\pi}(s, a)$ and $\rho_t^b(s, a)$ as the marginal density of $(S_t, A_t)$ at $(s, a) \in S \times A$ under the target policy $\pi$ and behavior policy $\pi^b$ respectively. Define the probability ratio function $w_t^{\pi}$ and Q-function as $w^{\pi}_t(s,a) = ho^{\pi}_t(s,a)/ ho^b_t(s,a), \qquad Q^{\pi}_t$ FQE: Recursively apply a regression technique to learn $Q_T^{\pi}, Q_{T-1}^{\pi}, \ldots, Q_1^{\pi}$ in a sequential and backward order. Let $\hat{Q}_{T+1}^{\pi} = 0$ , and for $t = T, T - 1, \ldots, 1$ , one can compute $\hat{Q}_t^{\pi} = \arg\min_{Q \in \mathcal{Q}^{(t)}} \frac{1}{n} \sum_{i=1}^n \left\{ Q(S_{i,t}, A_{i,t}) - \left[ R_{i,t} \right] \right\}$ and $\hat{\nu}(\pi) = \int_{(s,a)\in\mathcal{S}\times\mathcal{A}} \hat{Q}_1^{\pi}(s,a) \rho_1^{\pi}(s,a) d(s,a).$ ▶ One can let $Q^{(t)} = \{\phi_K(\cdot)^T \beta : \beta \in \mathbb{R}^{K|A|}\}$ , where $\phi_K$ are pre-specified features. We consider two scenarios: 1. K is fixed $\rightarrow$ parametric setting. 2. K is growing with n (and T) $\rightarrow$ nonparametric setting. Three key questions optimal convergence rate $(n^{-1/2})$ still achievable under nonparametric models of Q-functions?

- 2. How does the convergence rate depend on the growing horizon T?
- 3. What is the role of the probability ratio functions  $w_t^{\pi}$  in improving the convergence rate for FQE estimators?

#### Comparison on the error bounds

Table 1: Comparison on the error bound for the first-order term in existing works.  $\kappa$  is defined as an overlap constant.  $\tilde{\kappa}$  is the upper bound for the probability ratio functions; D is the dimension of space and action. d is the intrinsic dimension of the state-action space. Some logarithmic orders are omitted in the error bounds.

Work	PARAMETRIC?	Regu
Yin & Wang (2020)	$\checkmark$	
Duan et al. $(2020)$	$\checkmark$	
Zhang et al. $(2022)$		Dif
Nguyen-Tang et al. (2021)	×	
JI ET AL. $(2022)$	×	
Our Work FOR PARAMETRIC SETTING		
<b>Our Work</b> for nonparametric setting	G X	

### **Connection with MIS estimators**

MIS estimator under the tabular setting is shown to have an error bound that has a linear dependence on the horizon (Yin & Wang (2020)). There is an equivalence between the FQE estimator and MIS estimator when adopting linear modeling of Q-functions (Duan et al. (2020)). Is linear dependence on the horizon for more general linear modeling (with potentially continuous state space) achievable for FQE estimators?

# A Fine-grained Analysis of Fitted Q-evaluation: Beyond Parametric Models

<sup>1</sup>University of Texas at Dallas <sup>2</sup> The George Washington University <sup>3</sup> Texas A&M University

In reinforcement learning (RL), off-policy evaluation (OPE) is an important topic that focuses on estimating the expected total reward of a target policy

> We delve deeply into the analysis of FQE estimators within the framework of a finite-horizon, time-inhomogeneous Markov Decision Process (MDP).

$$\mathbb{E}^{\pi}(s,a) = \mathbb{E}^{\pi}\left[\sum_{t'=t}^{T} R_t \mid S_t = s, A_t = a\right].$$

$$_{t} + \sum_{a' \in \mathcal{A}_{t+1}} \pi_{t}(a \mid S_{i,t+1}) \hat{Q}_{t+1}^{\pi}(S_{i,t+1}, a') \bigg] \bigg\}^{2}$$

. For the fixed horizon T, how does the convergence rate depend on the number of episodes n given the completeness assumption for Q-functions? Is the

JLARITY ON $Q$	Error Bound
TABULAR	$\mathcal{O}(T \tilde{\kappa} \sqrt{1/n})$
LINEAR	$\mathcal{O}(T^2\sqrt{\kappa/n})$
FERENTIABLE	$\mathcal{O}(T^2\sqrt{\kappa/n})$
Besov	$\mathcal{O}(T^{2-lpha/(2lpha+2D)}\tilde{\kappa}n^{-lpha/(2lpha+2D)})$
Besov	$\mathcal{O}(T^2\kappa n^{-lpha/(2lpha+d)})$
LINEAR	$\mathcal{O}(T^{1.5}\sqrt{\kappa/n})$ $\mathcal{O}(T \tilde{\kappa} \sqrt{1/n})$ when $w_t^{\pi}$ are linear
Hölder	$\mathcal{O}(T^{1.5}\sqrt{\kappa/n})$ when $Q_t^{\pi}$ are smooth enough $\mathcal{O}(T\tilde{\kappa}\sqrt{1/n})$ when $w_t^{\pi}$ are Hölder

Jiayi Wang <sup>1</sup> Zhengling Qi <sup>2</sup> Raymond K. W. Wong <sup>3</sup>

### Parametric setting

Define  $(\mathcal{P}_t^{\pi} f)(s, a) = \mathbb{E}\left\{\sum_{a'} \pi_t(a' \mid S_{t+1}) f(S_{t+1}, a') \mid S_t = s, A_t = a\right\}, \ \kappa := \frac{1}{T} \sum_{t=1}^T \sup_{f \in \mathcal{O}(t)} [\mathcal{E}_t^{\pi} f]^2 / \|f\|_{\mathcal{L}_2}^2.$ 

- ► Assumption 1:  $\mathbb{E}\{R_t \mid S_t = \cdot, A_t = \cdot\} \in \mathcal{Q}^{(t)}$ , for t = 1, ..., T. For every  $q \in \mathcal{Q}^{(t+1)}$ , we have  $\mathcal{P}_t^{\pi} q \in \mathcal{Q}^{(t)}$ . **Theorem 1**:  $Q^{(t)} = \left\{ \phi_K(\cdot, \cdot)^{\mathsf{T}} \beta : \beta \in \mathbb{R}^{K|\mathcal{A}|} \right\}$  for some pre-specified feature  $\phi_K$  and K is a fixed constant, under Assumption 1 and some technical
- conditions, if  $T = \mathcal{O}([n/(\log n \log T)]^{1/2})$ , we have

$$|\hat{\nu}(\pi) - \nu(\pi)| = \mathcal{O}_{\mathrm{p}}\left(\sqrt{\frac{T^{3}\kappa}{n}} + T^{3}\frac{\log n \log T}{n}\right).$$

- $\triangleright$  Compared with the bound in Duan et al. (2020), our first order term has an order of  $T^{1.5}/\sqrt{n}$ . We achieve a sharper horizon dependence by exploiting the fact that the variance of the first order term can be decomposed as a sum of T individual expectations of the conditional variance. ► Assumption 2:  $w_t^{\pi} \in \left\{ \phi_K(\cdot, \cdot)^{\mathsf{T}} \beta : \beta \in \mathbb{R}^{K|\mathcal{A}|} \right\}, t = 1, \ldots, T.$
- **Theorem 2**: Under conditions listed in Theorem 1, if we futher assume Assumption 2, we have  $|\hat{\nu}(\pi) - \nu(\pi)| = \mathcal{O}_{\mathrm{p}}\left(T\sqrt{\frac{\kappa}{n}} + T^{3}\frac{\log n \log T}{n}\right)$
- > Theorem 2 shows that with an additional realizability assumption (Assumption 2) on the probability ratio functions, the convergence rate of the error will depend *linearly* with respect to the horizon T in the first-order term. This is a significant improvement in horizon dependence over the existing literature on the setting of using linear function approximation.

#### Nonparametric setting: a slower rate

Define the projection  $\Pi_t$  such that  $\Pi_t g(s, a) = \phi_K(s, a)^{\mathsf{T}}(\Sigma_t)^{-1} \mathbb{E} [\phi_K(S_t, A_t)g(S_t, A_t)].$ **Assumption 3:** For every  $a \in A$  and t = 1, ..., T,  $\{q(\cdot, a) : q \in Q^{(t)}, \|q\|_{\infty} \leq 1\}$  is a subset of Hölder space  $\Lambda_{\infty}(p, L)$  with constants p > d/2 and

- L > 0.
- **Assumption 4:** There exists a constant  $\beta_Q > 1/2$  (independent of T) such that  $\sup_{q \in Q^{(t)}(1)} \|q \Pi_t q\|_{\infty} \lesssim K^{-\beta_Q}$  for t = 1, ..., T.
- **Theorem 3:** Under Assumption 1, 3, 4 and some technical conditions, if we further assume that  $K = O(\min\{\sqrt{n/(\log n \log T)}, n/(T^2 \log n \log T)\})$ ,  $T = \mathcal{O}(K^{\beta_Q})$ , then we have

$$|\hat{\nu}(\pi) - \nu(\pi)| = \mathcal{O}_{p}\left(T^{2}K^{-\beta_{Q}} + \sqrt{\frac{T^{3}\kappa}{n}} + \frac{T^{3}K\log n\log T}{n}\right).$$

**Corollary 1:** Under conditions listed in Theorem 3, we further assume that T is bounded. (i) If  $1/2 < \beta_Q \leq 1$ , then by taking  $K \asymp \sqrt{n}/(\log n \log T)$ , we have

$$|\hat{\nu}(\pi) - \nu(\pi)| = \mathcal{O}_{\mathrm{p}}(n^{-\beta_Q/2})$$

(ii) If  $\beta_Q > 1$ , then by taking  $K \asymp (n/(\log n))^{1/(1+\beta_Q)}$ , we have

$$|\hat{\nu}(\pi) - \nu(\pi)| = \mathcal{O}_{\mathrm{p}}(n^{-1/2})$$

- $\triangleright$  When  $\beta_Q$  is large enough, i.e., Q functions are smooth enough, we can achieve the optimal convergence rate  $n^{-1/2}$ .
- $\blacktriangleright$  When  $1/2 < \beta_Q \le 1$ , by choosing K appropriately, our bound is faster than the optimal convergence rate  $n^{-\beta_Q/(1+2\beta_Q)}$  for nonparametrically estimating the Q-functions

#### Nonparametric setting: a faster rate with realizability on ratio function

- **Assumption 5:** There
- **Theorem 4:** Under con

exists a constant 
$$\beta_w > 1/2$$
 such that  $\sup_t ||w_t^{\pi} - \Pi_t w_t^{\pi}||_{\infty} \lesssim K^{-\beta_w}$  for  $t = 1, ..., T$ .  
additions listed in Theorem 3, we further assume Assumption 5, we have  
 $\hat{\nu}(\pi) - \nu(\pi)| = \mathcal{O}\left(\frac{T}{\sqrt{n}} + T^2 K^{-\beta_Q - \beta_w} + T^3 K^{-2\beta_Q} + T^3 K^{-\beta_Q} \sqrt{\frac{K \log n \log T}{n}} + \frac{T^3 K \log n \log T}{n}\right).$ 
Inditions listed in Theorem 4, we further assume  $\beta_Q = \beta_w = \beta > 1/2$ ,  $T \log T = \mathcal{O}\left((n/\log n)^{\beta/(1+2\beta)}\right)$ , by taking the optimal  $K \simeq (\pi/(\log n \log T))^{\frac{1}{1+2\beta}}$ 

**Corollary 2:** Under con order of K such that

$$\begin{split} \sup_{t} \|w_{t}^{\pi} - \Pi_{t} w_{t}^{\pi}\|_{\infty} &\lesssim K^{-\beta_{w}} \text{ for } t = 1, \dots, T. \\ \text{r assume Assumption 5, we have} \\ -\beta_{w} + T^{3} K^{-2\beta_{Q}} + T^{3} K^{-\beta_{Q}} \sqrt{\frac{K \log n \log T}{n}} + \frac{T^{3} K \log n \log T}{n} \Big). \\ \text{er assume } \beta_{Q} = \beta_{w} = \beta > 1/2, \ T \log T = \mathcal{O}\left((n/\log n)^{\beta/(1+2\beta)}\right), \text{ by taking the optimal} \\ K \approx \{n/(\log n \log T)\}^{\frac{1}{1+2\beta}}, \end{split}$$

we have 
$$|\hat{
u}(\pi) - 
u(\pi)| =$$

$$\begin{cases} \mathcal{O}_{\mathrm{p}}\left(\frac{T}{\sqrt{n}}\right), \text{if } T = \mathcal{O}\left(n^{\frac{2\beta-1}{4(1+2\beta)}}(\log n)^{\frac{-\beta}{1+2\beta}}\right), \\ \mathcal{O}_{\mathrm{p}}\left(T^{3}\left(\frac{n}{\log n\log T}\right)^{\frac{-2\beta}{1+2\beta}}\right), \text{otherwise.} \end{cases}$$

If the number of horizon T is bounded, we can achieve the optimal convergence rate  $(n^{-1/2})$  for  $|\hat{\nu}(\pi) - \nu(\pi)|$  even though we estimate Q functions nonparametricly. Compared to Corollary 1, we do not require  $\beta_Q > 1$  to achieve such optimal convergence rate. ln the scenario where T grows relatively slowly compared to n (case 1), the convergence exhibits a  $n^{-1/2}$  dependence with respect to n, with a linear dependence on the horizon. To the best of our knowledge, this convergence rate aligns with the best-known rate for FQE in tabular settings Yin & Wang (2020) (necessarily parametric), despite our analysis is conducted under a much more challenging nonparametric setting.



 $\log n$