

# Subsample Ridge Ensembles: Equivalences and Generalized Cross-Validation

Jin-Hong Du<sup>1\*</sup> Pratik Patil<sup>2\*</sup> Arun Kumar Kuchibhotla<sup>1</sup>

<sup>1</sup>Department of Statistics and Data Science, Carnegie Mellon University

<sup>2</sup>Department of Statistics, University of California, Berkeley

\*equal contribution

July 2023

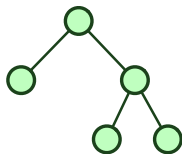
**Carnegie  
Mellon  
University**

**Berkeley**  
UNIVERSITY OF CALIFORNIA

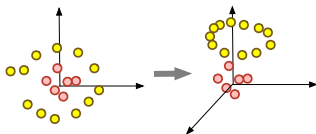
# Over-parameterization and regularization

- ▶ In the big data era, the success of machine learning and deep learning methods typically have much more parameters than the training samples.

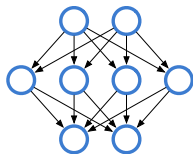
Random forest



Kernel method



Neural network



- ▶ Optimizing such over-parameterized models requires different types of regularization.

# Explicit and implicit regularization

implicit regularization



explicit regularization

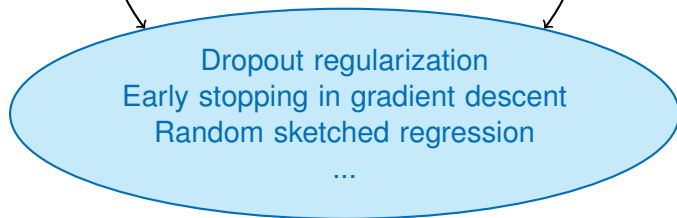


# Explicit and implicit regularization

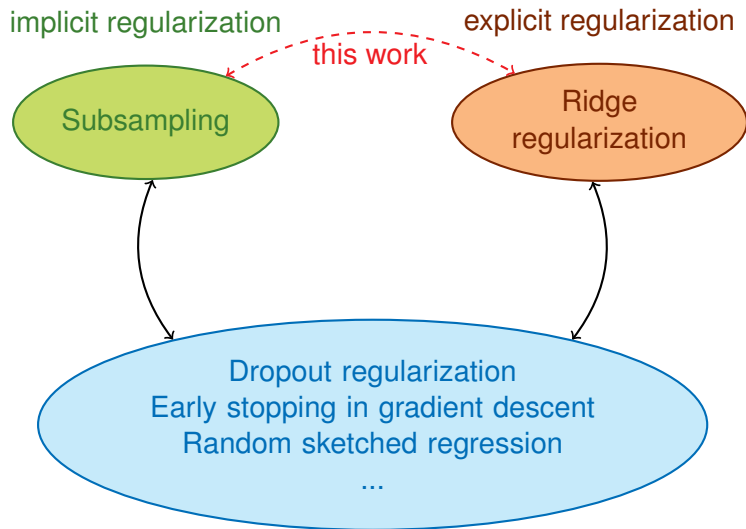
implicit regularization



explicit regularization



# Explicit and implicit regularization



# Ridge ensembles

- **Ridge estimator:** Let  $\mathcal{D}_n = \{(\mathbf{x}_j, y_j) \in \mathbb{R}^p \times \mathbb{R} : j \in [n]\}$  denote a dataset. The ridge estimator fitted on subsampled dataset  $\mathcal{D}_I$  with  $I \subseteq [n], |I| = k$  is defined as:

$$\hat{\beta}_k^\lambda(\mathcal{D}_I) = \operatorname{argmin}_{\beta \in \mathbb{R}^p} \frac{1}{k} \sum_{j \in I} (y_j - \mathbf{x}_j^\top \beta)^2 + \lambda \|\beta\|_2^2.$$

# Ridge ensembles

- ▶ **Ridge estimator:** Let  $\mathcal{D}_n = \{(\mathbf{x}_j, y_j) \in \mathbb{R}^p \times \mathbb{R} : j \in [n]\}$  denote a dataset. The ridge estimator fitted on subsampled dataset  $\mathcal{D}_I$  with  $I \subseteq [n], |I| = k$  is defined as:

$$\hat{\beta}_k^\lambda(\mathcal{D}_I) = \operatorname{argmin}_{\beta \in \mathbb{R}^p} \frac{1}{k} \sum_{j \in I} (y_j - \mathbf{x}_j^\top \beta)^2 + \lambda \|\beta\|_2^2.$$

- ▶ **Ensemble ridge estimator:**

$$\tilde{\beta}_{k,M}^\lambda(\mathcal{D}_n; \{I_\ell\}_{\ell=1}^M) := \frac{1}{M} \sum_{\ell \in [M]} \hat{\beta}_k^\lambda(\mathcal{D}_{I_\ell}),$$

with  $I_1, \dots, I_M \sim \mathcal{I}_k := \{\{i_1, \dots, i_k\} : 1 \leq i_1 < \dots < i_k \leq n\}$ . The *full-ensemble* ridge estimator is defined by letting  $M \rightarrow \infty$ .

# Prediction risk

**Conditional prediction risk:** The goal is to quantify and estimate the prediction risk:

$$R_{k,M}^\lambda := \mathbb{E}_{(x,y)}[(y - \mathbf{x}^\top \tilde{\beta}_{k,M}^\lambda)^2 \mid \mathcal{D}_n, \{I_\ell\}_{\ell=1}^M], \quad (1)$$

under proportional asymptotics where  $n, p, k \rightarrow \infty$ ,  $p/n \rightarrow \phi$  and  $p/k \rightarrow \phi_s$ . Here,  $\phi$  and  $\phi_s$  are the *data* and *subsample* aspect ratios, respectively.



# Prediction risk

**Conditional prediction risk:** The goal is to quantify and estimate the prediction risk:

$$R_{k,M}^\lambda := \mathbb{E}_{(x,y)}[(y - \mathbf{x}^\top \tilde{\beta}_{k,M}^\lambda)^2 \mid \mathcal{D}_n, \{I_\ell\}_{\ell=1}^M], \quad (1)$$

under proportional asymptotics where  $n, p, k \rightarrow \infty$ ,  $p/n \rightarrow \phi$  and  $p/k \rightarrow \phi_s$ . Here,  $\phi$  and  $\phi_s$  are the *data* and *subsample* aspect ratios, respectively.

Focusing on subsample ridge ensemble, we aim to answer:

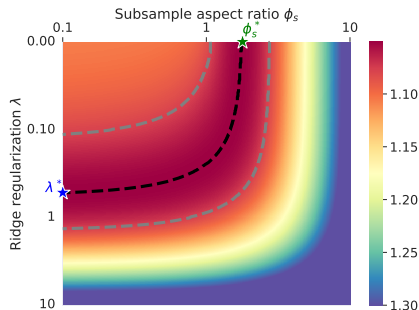
- (1) What is the role and relationship between implicit *subsampling* and explicit *ridge* regularization with regard to prediction risk?
- (2) How to tune the subsample aspect ratio  $\phi_s$  and the ridge penalty  $\lambda$  to minimize the prediction risk?

# Risk equivalence

- ▶ As  $p/n \rightarrow \phi$  and  $p/k \rightarrow \phi_s$ , the prediction risk in the full ensemble ( $M = \infty$ ) converges:

$$R_{k,\infty}^\lambda \xrightarrow{\text{a.s.}} \mathcal{R}_\infty^\lambda(\phi, \phi_s).$$

- ▶ For  $\phi = 0.1$ , the risk profile as a function of  $(\lambda, \phi_s)$  is shown in the figure in the log-log scale.

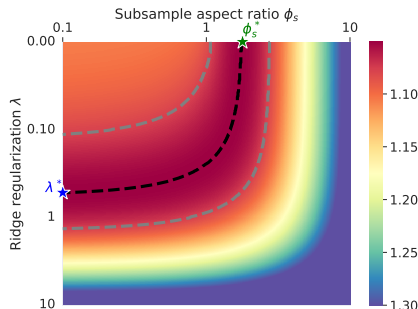


# Risk equivalence

- ▶ As  $p/n \rightarrow \phi$  and  $p/k \rightarrow \phi_s$ , the prediction risk in the full ensemble ( $M = \infty$ ) converges:

$$R_{k,\infty}^\lambda \xrightarrow{\text{a.s.}} \mathcal{R}_\infty^\lambda(\phi, \phi_s).$$

- ▶ For  $\phi = 0.1$ , the risk profile as a function of  $(\lambda, \phi_s)$  is shown in the figure in the log-log scale.
- ▶ Risk equivalence (Theorem 2.3):



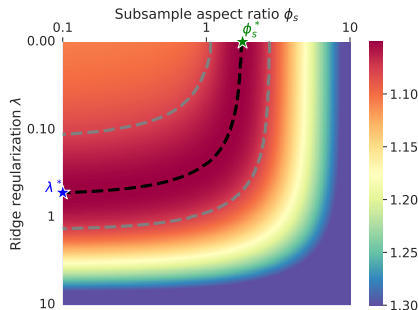
$$\underbrace{\min_{\phi_s \geq \phi} \mathcal{R}_\infty^0(\phi, \phi_s)}_{\text{opt. ridgeless ensemble}} = \underbrace{\min_{\lambda \geq 0} \mathcal{R}_\infty^\lambda(\phi, \phi)}_{\text{opt. ridge predictor}} = \underbrace{\min_{\substack{\phi_s \geq \phi, \\ \lambda \geq 0}} \mathcal{R}_\infty^\lambda(\phi, \phi_s)}_{\text{opt. ridge ensemble}}.$$

# Risk equivalence

- ▶ As  $p/n \rightarrow \phi$  and  $p/k \rightarrow \phi_s$ , the prediction risk in the full ensemble ( $M = \infty$ ) converges:

$$R_{k,\infty}^\lambda \xrightarrow{\text{a.s.}} \mathcal{R}_\infty^\lambda(\phi, \phi_s).$$

- ▶ For  $\phi = 0.1$ , the risk profile as a function of  $(\lambda, \phi_s)$  is shown in the figure in the log-log scale.
- ▶ Implication: the implicit regularization provided by the subsample ensemble (a larger  $\phi_s$ , or a smaller  $k$ ) amounts to adding more explicit ridge regularization (a larger  $\lambda$ ).



# Generalized cross-validation for ridge ensembles

- ▶ Beyond quantitative analysis, how can one pick  $(\lambda, \phi_s)$  to minimize the prediction risk?

# Generalized cross-validation for ridge ensembles

- ▶ Beyond quantitative analysis, how can one pick  $(\lambda, \phi_s)$  to minimize the prediction risk?
- ▶ For ordinary ridge ( $M = 1$  or  $k = n$ ), the **generalized cross-validation (GCV)** estimator is known to be consistent.

# Generalized cross-validation for ridge ensembles

- ▶ Beyond quantitative analysis, how can one pick  $(\lambda, \phi_s)$  to minimize the prediction risk?
- ▶ For ordinary ridge ( $M = 1$  or  $k = n$ ), the **generalized cross-validation (GCV)** estimator is known to be consistent.
- ▶ For general  $M$ , the GCV estimator is defined as

$$\text{gcv}_{k,M}^\lambda = \frac{T_{k,M}^\lambda}{D_{k,M}^\lambda} \quad \leftarrow \begin{array}{l} \text{training error} \\ \text{degree of freedom correction} \end{array}$$

# Generalized cross-validation for ridge ensembles

- ▶ Beyond quantitative analysis, how can one pick  $(\lambda, \phi_s)$  to minimize the prediction risk?
- ▶ For ordinary ridge ( $M = 1$  or  $k = n$ ), the **generalized cross-validation (GCV)** estimator is known to be consistent.
- ▶ For general  $M$ , the GCV estimator is defined as

$$\text{gcv}_{k,M}^\lambda = \frac{T_{k,M}^\lambda}{D_{k,M}^\lambda} = \frac{\frac{1}{|I_{1:M}|} \sum_{i \in I_{1:M}} (y_i - \mathbf{x}_i^\top \tilde{\beta}_{k,M}^\lambda)^2}{(1 - |I_{1:M}|^{-1} \text{tr}(\mathbf{S}_{k,M}^\lambda))^2},$$

where  $\mathbf{S}_{k,M}^\lambda = \frac{1}{M} \sum_{\ell=1}^M \mathbf{X}_{I_\ell} (\mathbf{X}_{I_\ell}^\top \mathbf{X}_{I_\ell} / k + \lambda \mathbf{I}_p)^+ \mathbf{X}_{I_\ell}^\top / k$  is the smoothing matrix that represents the degree of freedom.



# Generalized cross-validation for ridge ensembles

- ▶ Beyond quantitative analysis, how can one pick  $(\lambda, \phi_s)$  to minimize the prediction risk?
- ▶ For ordinary ridge ( $M = 1$  or  $k = n$ ), the **generalized cross-validation (GCV)** estimator is known to be consistent.
- ▶ For general  $M$ , the GCV estimator is defined as

$$\text{gcv}_{k,M}^\lambda = \frac{T_{k,M}^\lambda}{D_{k,M}^\lambda} = \frac{\frac{1}{|I_{1:M}|} \sum_{i \in I_{1:M}} (y_i - \mathbf{x}_i^\top \tilde{\beta}_{k,M}^\lambda)^2}{(1 - |I_{1:M}|^{-1} \text{tr}(\mathbf{S}_{k,M}^\lambda))^2},$$

where  $\mathbf{S}_{k,M}^\lambda = \frac{1}{M} \sum_{\ell=1}^M \mathbf{X}_{I_\ell} (\mathbf{X}_{I_\ell}^\top \mathbf{X}_{I_\ell} / k + \lambda \mathbf{I}_p)^+ \mathbf{X}_{I_\ell}^\top / k$  is the smoothing matrix that represents the degree of freedom.

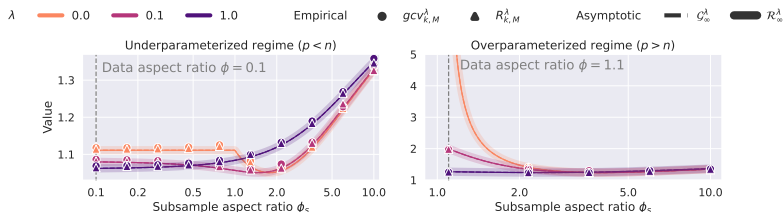
- ▶ The GCV for full ensemble is defined by letting  $M$  tend to infinity.

# Uniform consistency of GCV for full-ensemble ridge

- ▶ (Theorem 3.1, informal) For all  $\lambda \geq 0$ , we have

$$\max_{k \in \mathcal{K}_n} |\text{gcv}_{k, \infty}^\lambda - R_{k, \infty}^\lambda| \xrightarrow{\text{a.s.}} 0.$$

- ▶ This allows selecting the optimal ensemble and subsample sizes in a data-dependent manner:



Coupled with the risk equivalence result, it suffices to fix  $\lambda$  and only tune the subsample size  $k$  or subsample aspect ratio  $\phi_s$ .

# Inconsistency on finite ensembles

- ▶ (Proposition 3.3, informal) For ensemble size  $M = 2$ , ridge penalty  $\lambda = 0$ , and any  $\phi \in (0, \infty)$ ,

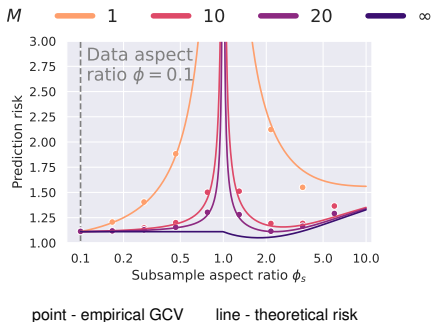
$$|\text{gcv}_{k,2}^0 - R_{k,2}^0| \not\rightarrow 0.$$

# Inconsistency on finite ensembles

- ▶ (Proposition 3.3, informal) For ensemble size  $M = 2$ , ridge penalty  $\lambda = 0$ , and any  $\phi \in (0, \infty)$ ,

$$|\text{gcv}_{k,2}^0 - R_{k,2}^0| \not\rightarrow 0.$$

- ▶ The bias scales as  $1/M$ , which is negligible for large  $M$ :



# Summary

- ▶ This work [1] reveals the connections between the *implicit regularization* induced by *subsampling* and *explicit ridge regularization* for subsample ridge ensembles.
- ▶ We establish the *uniform consistency* of GCV for full ridge ensembles.
- ▶ We show that GCV can be *inconsistent* even for ridge ensembles when  $M = 2$ .
- ▶ Future directions: bias correction of GCV for finite  $M$ ; extension to other metrics [2]; extension to other base predictors.

[1] Jin-Hong Du, Pratik Patil, and Arun Kumar Kuchibhotla. "Subsample Ridge Ensembles: Equivalences and Generalized Cross-Validation". In: *International Conference on Machine Learning* (2023)

[2] Pratik Patil and Jin-Hong Du. "Generalized equivalences between subsampling and ridge regularization". In: *arXiv preprint arXiv:2305.18496* (2023)