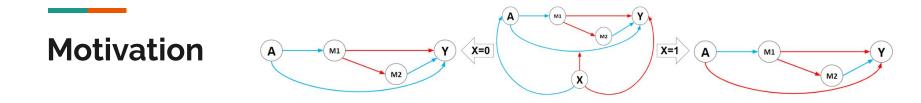
On Heterogeneous Treatment Effects in Heterogeneous Causal Graphs

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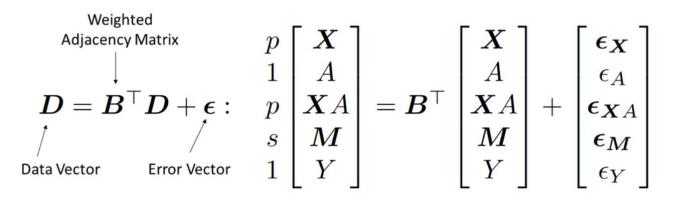
- Causal discovery, the task of discovering the causal relations between variables in a dataset, has interesting and important applications in many areas, such as epidemiology, medicine, economics, etc.
- Yet, due to the heterogeneity of individuals in response to different events/treatments, there may not exist a uniform causal graph for everyone. This implies the existence of heterogeneous causal graphs (HCGs).
- Very few studies have been conducted to investigate heterogeneous causal effects (HCEs) in **graphical contexts** due to the lack of statistical methods.

Contributions

- To our knowledge, this is the **first** work that considers heterogeneity in terms of causal graphs. We conceptualize HCGs, by **incorporating moderators** and their interaction directly into our model
- We compute the HCEs, directly from the HCG, to quantify the impact of the treatment and mediators on the outcome of interest given the moderators.
- We propose an **interactive structural learning algorithm** to learn the complex HCGs, estimate HCEs, and compute bootstrap confidence intervals for these estimates via a debiasing process.

Whole graph LSEM

- X is directly incorporated into model along with interaction
- B's **sparness** can be described using prior causal information
- No noise distribution assumed on the errors



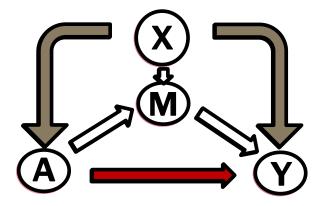
Whole graph to subgraph

- The do-operator is used to reveal the desired subgraph once the whole graph has been estimated
- This operation applies a value to the variable in question ignoring any parents.

$$D_{do(\boldsymbol{X}=\boldsymbol{x})} = \boldsymbol{B}_{do(\boldsymbol{X}=\boldsymbol{x})}^{\top} \boldsymbol{D}_{do(\boldsymbol{X}=\boldsymbol{x})} + \boldsymbol{\epsilon}' \rightarrow \\ \begin{bmatrix} A \\ \boldsymbol{M} \\ \boldsymbol{Y} \end{bmatrix} = \begin{bmatrix} \boldsymbol{\delta}_{\boldsymbol{X}} \boldsymbol{x} \\ \boldsymbol{B}_{\boldsymbol{X}}^{\top} \boldsymbol{x} \\ \boldsymbol{\gamma}_{\boldsymbol{X}} \boldsymbol{x} \end{bmatrix} + \boldsymbol{B}_{do(\boldsymbol{X}=\boldsymbol{x})}^{\top} \begin{bmatrix} A \\ \boldsymbol{M} \\ \boldsymbol{Y} \end{bmatrix} + \begin{bmatrix} \boldsymbol{\epsilon}_{A} \\ \boldsymbol{\epsilon}_{M} \\ \boldsymbol{\epsilon}_{Y} \end{bmatrix}$$

Heterogeneous Direct Effect of A on Y (HDE): The effect on Y that is due to A *and not mediated* by M given the baseline covariates, X

$$HDE(\boldsymbol{x}) = E[Y|do(A = a + 1, \boldsymbol{M} = \boldsymbol{m}^{(a)}), \boldsymbol{X} = \boldsymbol{x}] - E[Y|do(A = a), \boldsymbol{X} = \boldsymbol{x}]$$

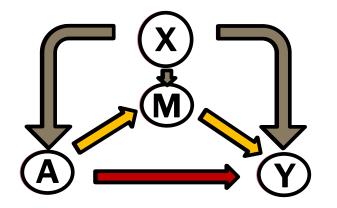


Heterogeneous Indirect Effect of A on Y (HIE): The effect on Y that is due to A *and mediated* by M given the baseline covariates, X

$$HIE(\mathbf{x}) = E[Y|do(A = a, M = m^{(a+1)}), \mathbf{X} = \mathbf{x}] - E[Y|do(A = a), \mathbf{X} = \mathbf{x}]$$

Heterogeneous Total Effect of A on Y (HTE): The total effect on Y that is due to A given the baseline covariates, X

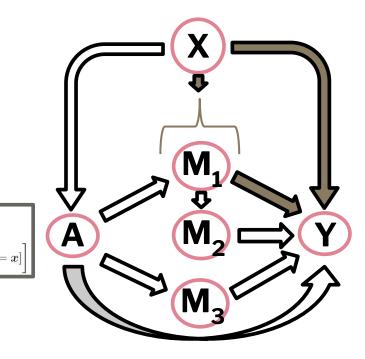
$$HTE(\boldsymbol{x}) = E[Y|do(A = a + 1), \boldsymbol{X} = \boldsymbol{x}] - E[Y|do(A = a), \boldsymbol{X} = \boldsymbol{x}]$$



Heterogeneous Direct Mediation Effect of M, on Y (HDM): The effect on Y that is due to M, and not mediated by other mediators given the baseline covariates, X

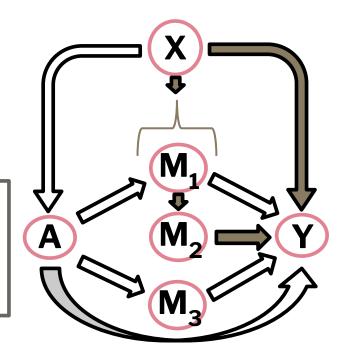
$$\begin{split} HDM_i(\bm{x}) &= \left\lfloor E[M_i| do(A = a + 1), \bm{X} = \bm{x}] - E[M_i| do(A = a), \bm{X} = \bm{x}] \right\rfloor \\ &\times \left[E[Y| do(A = a, M_i = m_i^{(a)} + 1, \bm{\Omega}_i = \bm{o}_i^{(a)}), \bm{X} = \bm{x}] - E[Y| do(A = a), \bm{X} = \bm{x}] \right] \end{split}$$

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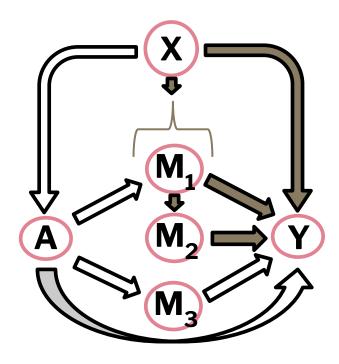
Heterogeneous Indirect Mediation Effect of M, **on Y** (**HIM**): The effect on Y that is due to M, *and mediated* by other mediators given the baseline covariates, X

$$HIM_{i}(\boldsymbol{x}) = \left[E[M_{i}|do(A = a + 1), \boldsymbol{X} = \boldsymbol{x}] - E[M_{i}|do(A = a), \boldsymbol{X} = \boldsymbol{x}] \right]$$
$$\times \left[E[Y|do(A = a, M_{i} = m_{i}^{(a)} + 1), \boldsymbol{X} = \boldsymbol{x}] - E[Y|do(A = a, M_{i} = m_{i}^{(a)} + 1, \boldsymbol{\Omega}_{i} = \boldsymbol{o}_{i}^{(a)}), \boldsymbol{X} = \boldsymbol{x}] \right]$$



Heterogeneous Total Mediation Effect of M, **on Y** (**HME**): The total effect on Y that is due to M_i given the baseline covariates, X

$$\begin{split} HME_i(\boldsymbol{x}) = & \left[E\{M_i | do(A = a + 1), \boldsymbol{X} = \boldsymbol{x}\} - E\{M_i | do(A = a), \boldsymbol{X} = \boldsymbol{x}\} \right] \\ & \times \left[E\{Y | do(M_i = m_i + 1), \boldsymbol{X} = \boldsymbol{x}\} - E\{Y | do(M_i = m_i), \boldsymbol{X} = \boldsymbol{x}\} \right] \end{split}$$



Theorem 4.1. Under assumptions (A1-A3) and the model described in Equation (1), we have: 1). $HDE(\mathbf{x}) = \gamma_A + \gamma_{\mathbf{X}A}\mathbf{x};$ 2). $HIE(\mathbf{x}) = \gamma_{\mathbf{M}}(\mathbf{I}_s - \mathbf{B}_{\mathbf{M}}^{\top})^{-1}(\boldsymbol{\beta}_A + \mathbf{B}_{\mathbf{X}A}^{\top}\mathbf{x});$ 3). $HTE(\mathbf{x}) = HDE(\mathbf{x}) + HIE(\mathbf{x}), \text{ where } \mathbf{I}_s \text{ is a } s \times s$ identity matrix and \mathbf{x} is the value of \mathbf{X} .

Theorem 4.2. Under assumptions (A1-A3) and Model (1), 1a). $HDM_i(\mathbf{x}) = \{\gamma_M\}_i\{(\mathbf{I}_s - \mathbf{B}_M^\top)^{-1}(\beta_A + \mathbf{B}_{\mathbf{X}A}^\top \mathbf{x})\}_i;$ 1b). $\sum_{i=1}^{s} HDM_i(\mathbf{x}) = HIE(\mathbf{x});$ 2). $HIM_i(\mathbf{x}) = HTM_i(\mathbf{x}) - HDM_i(\mathbf{x});$ 3). $HTM_i(\mathbf{x}) = HIE(\mathbf{x}) - HIE_{\mathbb{G}(-i)}(\mathbf{x}),$ where $\{\cdot\}_i$ is the ith element of a vector and $HIE_{\mathbb{G}(-i)}$ is the HIE under the causal graph $\mathbb{G}(-i)$ in which the ith mediator is removed from the original causal graph \mathbb{G} .

$$\boldsymbol{B}^{\top} = \begin{bmatrix} \boldsymbol{0}_{p \times p} & \boldsymbol{0}_{p \times 1} & \boldsymbol{0}_{p \times p} & \boldsymbol{0}_{p \times s} & \boldsymbol{0}_{p \times 1} \\ \boldsymbol{\delta}_{\boldsymbol{X}} & \boldsymbol{0} & \boldsymbol{0}_{1 \times p} & \boldsymbol{0}_{1 \times s} & \boldsymbol{0} \\ \boldsymbol{0}_{p \times p} & \boldsymbol{0}_{p \times 1} & \boldsymbol{0}_{p \times p} & \boldsymbol{0}_{p \times s} & \boldsymbol{0}_{p \times 1} \\ \boldsymbol{B}_{\boldsymbol{X}}^{\top} & \boldsymbol{\beta}_{\boldsymbol{A}} & \boldsymbol{B}_{\boldsymbol{X}\boldsymbol{A}}^{\top} & \boldsymbol{B}_{\boldsymbol{M}}^{\top} & \boldsymbol{0}_{s \times 1} \\ \boldsymbol{\gamma}_{\boldsymbol{X}} & \boldsymbol{\gamma}_{\boldsymbol{A}} & \boldsymbol{\gamma}_{\boldsymbol{X}\boldsymbol{A}} & \boldsymbol{\gamma}_{\boldsymbol{M}} & \boldsymbol{0} \end{bmatrix}$$

- Functional versions of both theorems exist
- Sparsity of B can be characterized

Structural Learning

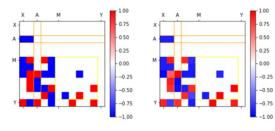
- Equation 1-4 describe the necessary sparsity of B
- h2 can be used alongside the acyclity constraint commonly used in score-based methods
 - 1. X has no parents, i.e., $g_1(B) = \sum_{j=1}^{p} \sum_{i=1}^{2p+s+2} |b_{i,j}| = 0;$ 3. Y has no descendants, i.e., $g_3(B) = \sum_{i=1}^{2p+s+2} |b_{2p+s+2,i}| = 0;$ 2. the order parents of A are X i.e., 4. the interaction X A also does not have parent
 - 2. the only parents of A are \boldsymbol{X} , i.e.,

4. the interaction XA also does not have parents, i.e.,

$$g_2(\mathbf{B}) = \sum_{i=p+1}^{2p+s+2} |b_{i,p+1}| = 0; \qquad \qquad g_4(\mathbf{B}) = \sum_{j=p+2}^{2p+1} \sum_{i=1}^{2p+s+2} |b_{i,j}| = 0$$

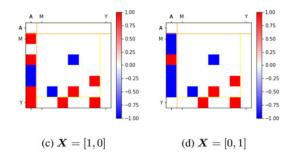
$$h_2(B) = \sum_{i=1}^4 g_i(B) = 0$$

Simulation Results









- In total there were 8 simulation scenarios and a real data study
- To the left and bottom are the results from scenario 3
- Overall, ISL makes accurate estimates and outperforms its competitors.

	FDR	TPR	SHD
NOTEARS	0.10(0.09)	0.94(0.08)	2.52(1.97)
DAG-GNN	0.03(0.05)	0.98(0.04)	0.70(1.05)
ISL	0.00(0.01)	1.00(0.01)	0.04(0.24)

Thank you for listening!