ARMS: Antithetic-REINFORCE-Multi-Sample Gradient for Binary Variables

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Motivation

Background

- Baseline gradients
- (Dis)ARM
- Copulas

ARMS

- Two sample estimator
- Multi sample estimator
- Antithetic Copulas

Goal: expectation based objectives of the form $\mathbb{E}_{p_{\phi}(x)}[f(x)]$

Can be found in many areas:

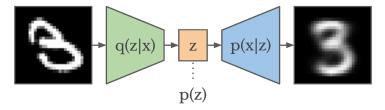
- Variational inference: $\mathbb{E}_{q_{\phi}(z|x)}[\ln p(x,z) \ln q_{\phi}(z|x)]$
- Reinforcement learning: $\mathbb{E}_{\pi_{\phi}(x)}[\sum_{t=0}^{T} \gamma^{t} r(s_{t}, a_{t})]$
- Finance (options pricing): $\mathbb{E}_{p_{s_0}(s_T)}[e^{-\gamma T} \max\{s_T K, 0\}]$
- Operations research (discrete queuing): $\mathbb{E}_{p_{\phi}(y_{1:T})} \left[\frac{\sum_{t=1}^{T} L_t(y_{1:t})}{\tau(y_{1:T})} \right]$
- Experimental design: $\mathbb{E}_{p_{\phi}(y)}[\mathbb{1}_{y < y_{\text{best}}}]$
- Et al. SDEs, GANs, bandits and online learning, econometrics, instrumental variables, counterfactual reasoning, ...

Discrete variational autoencoder

• Maximize evidence lower bound (ELBO):

$$\mathcal{L}_{\mathsf{ELBO}} = \mathbb{E}_{q_{\phi}(oldsymbol{z}|oldsymbol{x})}[\ln p_{ heta}(oldsymbol{z},oldsymbol{x}) - \ln q_{\phi}(oldsymbol{z}|oldsymbol{x})] \leq \ln p(oldsymbol{x})$$

- Equivalent to minimizing KL divergence KL(q(z|x)||p(z|x))
- $q(\pmb{z}|\pmb{x})$ is called an encoder, usually deep neural network $\pmb{x} o \phi$
- p(x|z) is a decoder, also neural network
- p(z) is a prior distribution
- In a discrete VAE $q(\boldsymbol{z}|\boldsymbol{x}) = \prod_{d=1}^{D} \text{Bern}(z_d|\sigma(\phi_d))$, with logits ϕ



- Why Monte Carlo gradients?
- Most expectations are too complicated to integrate
- Why the score function (REINFORCE) gradient?
- Most general, works with almost any distribution:

$$\nabla_{\phi} \mathbb{E}_{p_{\phi}(x)}[f(x)] = \mathbb{E}_{p_{\phi}(x)}[f(x)\nabla_{\phi} \ln p_{\phi}(x)]$$

• Unbiased: $E[g_{\mathsf{RF}}(x)] = E[f(x)\nabla_{\phi} \ln p_{\phi}(x)] = \nabla_{\phi}$

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- General REINFORCE: $f(z) \nabla \ln p_{\phi}(z)$. What if z is binary?
- Let $b \sim \text{Bern}(p)$, $p = \sigma(\phi)$.
- REINFORCE: $g_{RF} = f(b)(b p)$.
- What if we have *n* independent samples?
- LOORF (Leave One Out REINFORCE):

$$g_{\text{LOORF}} = \frac{1}{n-1} \sum_{i=1}^{n} \left(f(b_i) - \frac{1}{n} \sum_{j=1}^{n} f(b_j) \right) (b_i - p)$$

- What if we want to use antithetic (negatively correlated) pairs?
- If $u \sim \text{Unif}(0, 1)$, then $\mathbb{1}_{u < p} \sim \text{Bern}(p)$.
- ARM: $g_{\text{ARM}} = \left(f(\mathbb{1}_{u < p}) f(\mathbb{1}_{u > 1-p})\right) \left(\frac{1}{2} u\right).$
- Let $b = \mathbb{1}_{u < p}$, $b' = \mathbb{1}_{u > 1-p}$.
- DisARM/U2G noticed randomness can be integrated out:

$$g_{\text{DisARM}} = \frac{1}{2} (f(b) - f(b'))(b - b') \max(p, 1 - p)$$

- *n*-dimensional or multivariate probability distribution with uniform marginals.
- $\boldsymbol{u} = (u_1, ..., u_n) \sim C_n$, such that $\forall i : u_i \sim \text{Unif}(0, 1)$.
- How to create a copula?
- \bullet Start with some multivariate distribution $\mathcal{M}.$
- Calculate all marginal CDFs: $\forall i : F_{x_i}(x)$.
- Apply them to any sample: $\mathbf{x} = (x_1, ..., x_n) \sim \mathcal{M}$.
- $\boldsymbol{u} = (u_1, ..., u_n)$, where $u_i = F_{x_i}(x_i)$, is a copula sample.

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- *n* independent samples (LOORF) perform better than *n*/2 antithetic pairs (DisARM)
- Is it possible to combine the two approaches?
- Yes! If we can:
- Sample *n* antithetic variables
- Debias the estimator

• LOORF for n = 2? $b, b' \sim \text{Bern}(p)$

$$g_{2-\text{LOORF}} = \frac{1}{2} \Big(f(b) - f(b') \Big) (b - b')$$

• DisARM for $b = \mathbb{1}_{u < p}$, $b' = \mathbb{1}_{(1-u) < p}$

$$g_{\mathsf{DisARM}} = \frac{1}{2} \Big(f(b) - f(b') \Big) (b - b') \max(p, 1 - p)$$

Two extremes: no correlation ←→ minimal correlation (antithetic).
Can we generalize to an arbitrarily dependent Bernoulli pair? Yes!

- Let (b, b') ~ B₂(p) denote a sample from a bivariate Bernoulli distribution with correlation ρ = corr(b, b').
- An unbiased estimator is:

$$g_{\text{ARTS}} = rac{g_{2\text{LOORF}}}{1-
ho} = rac{1}{2} \Big(f(b) - f(b') \Big) (b-b') rac{1}{1-
ho}$$

- If $\rho = 0$, we obtain two sample LOORF.
- What is the lowest possible correlation for a Bernoulli pair b, b'?
- $\rho = -\min(\frac{p}{1-p}, \frac{1-p}{p})$, which results in DisARM.

• Key observation: LOORF is exactly the same as averaging all $\binom{n}{2}$ pairs!

$$g_{\text{LOORF}}(b_1, ..., b_n) = \frac{1}{n} \sum_{i=1}^n \left(f(b_i) - \frac{1}{n} \sum_{j=1}^n f(b_j) \right) \nabla_\phi \ln p(b_i)$$

= $\frac{1}{n(n-1)} \sum_{i \neq j} \frac{1}{2} \left(f(b_i) - f(b_j) \right) \left(\nabla_\phi \ln p(b_i) - \nabla_\phi \ln p(b_j) \right)$
= $\frac{1}{n(n-1)} \sum_{i \neq j} g_{2\text{-LOORF}}(b_i, b_j),$

• Idea: what if the debiasing term is identical for all pairs?

- Assume we have *n* Bernoulli variables, with $\rho = \operatorname{corr}(b_i, b_j)$, $\forall i \neq j$.
- An unbiased gradient is:

$$g_{\text{ARMS}} = rac{g_{\text{LOORF}}}{1-\rho} = rac{1}{n-1} \sum_{i=1}^{n} \left(f(b_i) - rac{1}{n} \sum_{j=1}^{n} f(b_j) \right) rac{b_i - p}{1-\rho}$$

- Last hurdle: how to sample *n* antithetic Bernoulli?
- If we have *n* antithetic uniform variables.
- Then $b_i = \mathbb{1}_{u_i < p}$ are antithetic Bernoulli.
- How to obtain *n* antithetic uniform variables? Copulas!
- (Sidenote: must also be able to calculate rho)

- Gaussian copula: numerical marginal and bivariate CDFs.
- Sample $(x_1, ..., x_n) \sim \mathcal{N}(0, \Sigma)$, with $\Sigma_{ii} = 1$, $\Sigma_{ij} = -1/(n-1)$.
- Let $u_i = \Phi(x_i)$, where $\Phi(x)$ is the standard Gaussian CDF.
- $(u_1, ..., u_n)$ is a copula with pairwise correlation close to to ρ_{min} .
- The correlation between to Bernoulli variables is: $\rho = \frac{\mathbb{E}[b_i b_j] p^2}{p(1-p)}$
- $E[b_i b_j] = P(b_i = b_j = 1) = P(u_i < p, u_j < p) = \Phi(p, p).$

- When $\alpha = 1$, we have both a analytical marginal and bivariate CDF.
- Sample $(d_1, ..., d_n) \sim \text{Dir}(1, ..., 1)$.
- Can alternatively do the following:

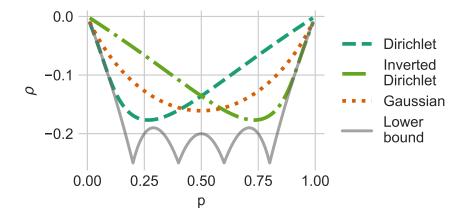
• 1.
$$v_i \sim \text{Unif}(0, 1), i = 1 \dots n$$

• 2.
$$d_i = \ln(v_i) / \sum_{j=1}^n \ln(v_j)$$

•
$$u_i = 1 - (1 - d_i)^{n-1}$$

• $\rho = \frac{\max(0, 2(1-p)^{\frac{1}{n-1}} - 1)^{n-1} - (1-p)^2}{p(1-p)}$

How good are the correlations?



- Sample *n* antithetic uniform variables (either Dirichlet or Gaussian)
- **2** Transform to Bernoulli $b_i = \mathbb{1}_{u_i < p}$ and calculate ρ
- Obtain an unbiased estimator ARMS:

$$g_{\text{ARMS}} = \frac{1}{n-1} \sum_{i=1}^{n} \left(f(b_i) - \frac{1}{n} \sum_{j=1}^{n} f(b_j) \right) \frac{b_i - p}{1 - \rho}$$

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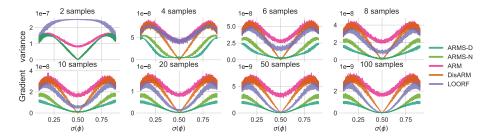
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Toy example

- Maximize: $\mathcal{E}(\phi) = \mathbb{E}_b[(b 0.499)^2], \quad b \sim \text{Bern}(\sigma(\phi))$
- Below: variance (all already unbiased) of each gradient as the function is maximized from $p_{init} = 0.05$ to $p_{end} = 0.95$.



Training ELBO

CAMPLES			ADMOD	ADMO M	LOODE	DraADM	DELAY
SAMPLES			ARMS-D	ARMS-N	LOORF	DISARM	RELAX
DYNAMIC MNIST	LINEAR	4	-112.13 ± 0.10	-111.96 ± 0.09	-112.32 ± 0.04	-113.26 ± 0.05	-112.98 ± 0.25
		6	-111.03 ± 0.02	-110.89 ± 0.07	-110.99 ± 0.07	-112.11 ± 0.03	-111.46 ± 0.06
		8	-110.30 ± 0.04	-110.62 ± 0.06	-110.42 ± 0.04	-111.78 ± 0.07	-110.58 ± 0.01
		10	-110.08 ± 0.05	-110.14 ± 0.09	-110.17 ± 0.04	-111.08 ± 0.11	-110.17 ± 0.09
	NONLINR	4	-98.65 ± 0.16	-98.97 ± 0.13	$\textbf{-98.62} \pm \textbf{0.05}$	-100.45 ± 0.16	-100.52 ± 0.08
		6	-98.53 ± 0.13	-97.87 ± 0.01	-98.14 ± 0.18	-99.28 ± 0.11	-99.17 ± 0.17
		8	-97.90 ± 0.12	-97.89 ± 0.10	-98.14 ± 0.21	-98.69 ± 0.21	-98.80 ± 0.02
		10	-97.64 ± 0.06	-97.32 ± 0.11	-97.50 ± 0.29	-98.62 ± 0.12	-98.69 ± 0.07
FASHION MNIST	LINEAR	4	-252.56 ± 0.11	-252.69 ± 0.06	-252.71 ± 0.09	-254.02 ± 0.05	-253.53 ± 0.06
		6	-251.94 ± 0.13	-251.73 ± 0.05	-252.03 ± 0.08	-252.97 ± 0.06	-252.31 ± 0.14
		8	-251.32 ± 0.11	-251.11 ± 0.23	-251.41 ± 0.10	-252.57 ± 0.05	-251.36 ± 0.08
		10	-251.29 ± 0.02	$\textbf{-251.08} \pm \textbf{0.08}$	-251.26 ± 0.03	-251.75 ± 0.21	$\textbf{-251.16} \pm \textbf{0.06}$
	NONLINR	4	-235.65 ± 0.12	-235.75 ± 0.06	-235.80 ± 0.07	-236.54 ± 0.06	-236.77 ± 0.03
		6	-235.47 ± 0.19	-235.36 ± 0.08	-235.70 ± 0.13	-235.94 ± 0.05	-236.20 ± 0.25
		8	-235.41 ± 0.10	-235.19 ± 0.14	-235.40 ± 0.13	-235.62 ± 0.16	-235.70 ± 0.14
		10	-235.18 ± 0.11	-235.32 ± 0.05	-235.59 ± 0.01	-235.60 ± 0.09	-235.46 ± 0.19
OMNIGLOT	Linear	4	-118.25 ± 0.08	-118.27 ± 0.05	-118.41 ± 0.07	-119.24 ± 0.17	-118.75 ± 0.08
		6	-117.62 ± 0.01	-117.62 ± 0.04	-117.75 ± 0.08	-118.47 ± 0.12	-117.90 ± 0.03
		8	-117.60 ± 0.05	-117.66 ± 0.12	-117.74 ± 0.10	-118.41 ± 0.10	-117.71 ± 0.02
		10	-117.03 ± 0.09	-116.99 ± 0.04	-117.21 ± 0.08	-117.70 ± 0.01	-117.13 ± 0.05
	NONLINR	4	-112.09 ± 0.27	$\textbf{-112.03} \pm \textbf{0.12}$	-112.20 ± 0.26	-113.24 ± 0.16	-114.08 ± 0.35
		6	-111.50 ± 0.06	-111.39 ± 0.10	-111.26 ± 0.15	-112.30 ± 0.05	-113.71 ± 0.13
		8	-110.91 ± 0.04	-111.01 ± 0.06	-110.85 ± 0.35	-111.82 ± 0.09	-113.64 ± 0.10
		10	-110.66 ± 0.05	-110.79 ± 0.26	-110.79 ± 0.20	-111.33 ± 0.19	-114.00 ± 0.10

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- ARMS generalizes *n* iid samples (LOORF) and two antithetic samples (DisARM)
- Future work: extension to categorical variables, IWAE bound with antithetic samples, ...

References

- Dong, Zhe, Andriy Mnih, and George Tucker. "DisARM: An Antithetic Gradient Estimator for Binary Latent Variables". In: *Advances in Neural Information Processing Systems 33.* 2020.
- Kool, Wouter, Herke van Hoof, and Max Welling. "Buy 4 REINFORCE Samples, Get a Baseline for Free!" In: *Workshop, Deep Reinforcement Learning Meets Structured Prediction, ICLR.* 2019.
- Mohamed, Shakir et al. "Monte Carlo Gradient Estimation in Machine Learning". In: J. Mach. Learn. Res. 21 (2020), 132:1–132:62.
- Salimans, Tim and David A Knowles. "On using control variates with stochastic approximation for variational bayes and its connection to stochastic linear regression". In: *arXiv preprint arXiv:1401.1022* (2014).
- Williams, Ronald J. "Simple statistical gradient-following algorithms for connectionist reinforcement learning". In: *Machine learning* 8.3-4 (1992), pp. 229–256.

Yin, Mingzhang et al. "Probabilistic Best Subset Selection by Gradient-Based Optimization". In: *arXiv e-prints* (2020).

Thank You! Questions Welcome!