# Subsample Ridge Ensembles: Equivalences and Generalized Cross-Validation

Jin-Hong Du<sup>1\*</sup> Pratik Patil<sup>2\*</sup> Arun Kumar Kuchibhotla<sup>1</sup>

<sup>1</sup>Department of Statistics and Data Science, Carnegie Mellon University <sup>2</sup>Department of Statistics, University of California, Berkeley equal contribution

July 2023





# Regularization

In the big data era, the success of machine learning and deep learning methods typically have much more parameters than the training samples.



 Optimizing such overparameterized models requires different types of regularization.

# Explicit and implicit regularization

#### implicit regularization



#### explicit regularization



# Explicit and implicit regularization



# Explicit and implicit regularization



# **Ridge ensembles**

► Ridge estimator: Let D<sub>n</sub> = {(x<sub>j</sub>, y<sub>j</sub>) ∈ ℝ<sup>p</sup> × ℝ : j ∈ [n]} denote a dataset. The ridge estimator fitted on subsampled dataset D<sub>I</sub> with I ⊆ [n], |I| = k is defined as:

$$\widehat{\boldsymbol{\beta}}_k^{\boldsymbol{\lambda}}(\mathcal{D}_I) = \operatorname*{argmin}_{\boldsymbol{\beta} \in \mathbb{R}^p} \frac{1}{k} \sum_{j \in I} (y_j - \boldsymbol{x}_j^\top \boldsymbol{\beta})^2 + \lambda \|\boldsymbol{\beta}\|_2^2$$

# **Ridge ensembles**

► Ridge estimator: Let D<sub>n</sub> = {(x<sub>j</sub>, y<sub>j</sub>) ∈ ℝ<sup>p</sup> × ℝ : j ∈ [n]} denote a dataset. The ridge estimator fitted on subsampled dataset D<sub>I</sub> with I ⊆ [n], |I| = k is defined as:

$$\widehat{\boldsymbol{\beta}}_k^{\boldsymbol{\lambda}}(\mathcal{D}_I) = \operatorname*{argmin}_{\boldsymbol{\beta} \in \mathbb{R}^p} \frac{1}{k} \sum_{j \in I} (y_j - \boldsymbol{x}_j^\top \boldsymbol{\beta})^2 + \lambda \|\boldsymbol{\beta}\|_2^2$$

Ensemble ridge estimator:

$$\widetilde{\boldsymbol{\beta}}_{k,\boldsymbol{M}}^{\lambda}(\mathcal{D}_n; \{I_\ell\}_{\ell=1}^{\boldsymbol{M}}) := \frac{1}{\boldsymbol{M}} \sum_{\ell \in [\boldsymbol{M}]} \widehat{\boldsymbol{\beta}}_k^{\lambda}(\mathcal{D}_{I_\ell}),$$

with  $I_1, \ldots, I_M \sim \mathcal{I}_k := \{\{i_1, \ldots, i_k\} : 1 \le i_1 < \ldots < i_k \le n\}$ . The *full-ensemble* ridge estimator is defined by letting  $M \to \infty$ .

# **Conditional prediction risk:** The goal is to quantify and estimate the prediction risk:

$$R_{k,M}^{\lambda} := \mathbb{E}_{(\boldsymbol{x},\boldsymbol{y})}[(\boldsymbol{y} - \boldsymbol{x}^{\top} \widetilde{\boldsymbol{\beta}}_{k,M}^{\lambda})^2 \mid \mathcal{D}_n, \{I_\ell\}_{\ell=1}^M],$$
(1)

under proportional asymptotics where  $n, p, k \to \infty$ ,  $p/n \to \phi$  and  $p/k \to \phi_s$ . Here,  $\phi$  and  $\phi_s$  are the data and subsample aspect ratios, respectively.

► As  $p/n \rightarrow \phi$  and  $p/k \rightarrow \phi_s$ , the prediction risk in the full ensemble ( $M = \infty$ ) converges:

$$R_{k,\infty}^{\lambda} \xrightarrow{\text{a.s.}} \mathscr{R}_{k,\infty}^{\lambda}(\phi,\phi_s).$$

For φ = 1, the risk profile as a function of (λ, φ<sub>s</sub>) is shown in the figure in the log-log scale.



► As  $p/n \rightarrow \phi$  and  $p/k \rightarrow \phi_s$ , the prediction risk in the full ensemble ( $M = \infty$ ) converges:

$$R_{k,\infty}^{\lambda} \xrightarrow{\text{a.s.}} \mathscr{R}_{k,\infty}^{\lambda}(\phi,\phi_s).$$

For φ = 1, the risk profile as a function of (λ, φ<sub>s</sub>) is shown in the figure in the log-log scale.







► As  $p/n \rightarrow \phi$  and  $p/k \rightarrow \phi_s$ , the prediction risk in the full ensemble ( $M = \infty$ ) converges:

 $R_{k,\infty}^{\lambda} \xrightarrow{\text{a.s.}} \mathscr{R}_{k,\infty}^{\lambda}(\phi,\phi_s).$ 

For φ = 1, the risk profile as a function of (λ, φ<sub>s</sub>) is shown in the figure in the log-log scale.



Implication: the implicit regularization provided by the subsample ensemble (a larger φ<sub>s</sub>, or a smaller k) amounts to adding more explicit ridge regularization (a larger λ).

Beyond quantitative analysis, how can one pick  $(\lambda, \phi_s)$  to minimize the prediction risk?

- Beyond quantitative analysis, how can one pick  $(\lambda, \phi_s)$  to minimize the prediction risk?
- For ordinary ridge (M = 1 or k = n), the generalized cross-validation (GCV) estimator is known to be consistent.

- Beyond quantitative analysis, how can one pick  $(\lambda, \phi_s)$  to minimize the prediction risk?
- For ordinary ridge (M = 1 or k = n), the generalized cross-validation (GCV) estimator is known to be consistent.
- For general M, the GCV estimator is defined as

$$\operatorname{gcv}_{k,M}^{\lambda} = \frac{T_{k,M}^{\lambda}}{D_{k,M}^{\lambda}} \quad \longleftarrow \quad \begin{array}{c} \operatorname{training \ error} \\ \operatorname{degree \ of \ freedom \ correction} \end{array}$$

- Beyond quantitative analysis, how can one pick  $(\lambda, \phi_s)$  to minimize the prediction risk?
- For ordinary ridge (M = 1 or k = n), the generalized cross-validation (GCV) estimator is known to be consistent.
- ▶ For general *M*, the GCV estimator is defined as

$$\mathsf{gcv}_{k,M}^{\lambda} = \frac{T_{k,M}^{\lambda}}{D_{k,M}^{\lambda}} = \frac{\frac{1}{|I_{1:M}|} \sum_{i \in I_{1:M}} (y_i - \boldsymbol{x}_i^{\top} \widetilde{\boldsymbol{\beta}}_{k,M}^{\lambda})^2}{(1 - |I_{1:M}|^{-1} \operatorname{tr}(\boldsymbol{S}_{k,M}^{\lambda}))^2},$$

where  $S_{k,M}^{\lambda} = \frac{1}{M} \sum_{\ell=1}^{M} X_{I_{\ell}} (X_{I_{\ell}}^{\top} X_{I_{\ell}}/k + \lambda I_p)^+ X_{I_{\ell}}^{\top}/k$  is the smoothing matrix that represents the degree of freedom.

- Beyond quantitative analysis, how can one pick  $(\lambda, \phi_s)$  to minimize the prediction risk?
- For ordinary ridge (M = 1 or k = n), the generalized cross-validation (GCV) estimator is known to be consistent.
- ▶ For general *M*, the GCV estimator is defined as

$$\mathsf{gcv}_{k,M}^{\lambda} = \frac{T_{k,M}^{\lambda}}{D_{k,M}^{\lambda}} = \frac{\frac{1}{|I_{1:M}|} \sum_{i \in I_{1:M}} (y_i - \boldsymbol{x}_i^\top \widetilde{\boldsymbol{\beta}}_{k,M}^{\lambda})^2}{(1 - |I_{1:M}|^{-1} \operatorname{tr}(\boldsymbol{S}_{k,M}^{\lambda}))^2},$$

where  $S_{k,M}^{\lambda} = \frac{1}{M} \sum_{\ell=1}^{M} X_{I_{\ell}} (X_{I_{\ell}}^{\top} X_{I_{\ell}}/k + \lambda I_p)^+ X_{I_{\ell}}^{\top}/k$  is the smoothing matrix that represents the degree of freedom.

▶ The GCV for full ensemble is defined by letting *M* tend to infinity.

# Uniform consistency of GCV for full-ensemble ridge

• (Theorem 3.1, informal) For all  $\lambda \ge 0$ , we have

$$\max_{k\in\mathcal{K}_n} |\mathsf{gcv}_{k,\infty}^\lambda - R_{k,\infty}^\lambda| \xrightarrow{\mathsf{a.s.}} 0.$$

This allows selecting the optimal ensemble and subsample sizes in a data-dependent manner:



Coupled with the risk equivalence result, it suffices to fix  $\lambda$  and only tune the subsample size *k* or subsample aspect ratio  $\phi_s$ .

# Inconsistency on finite ensembles

Proposition 3.3, informal) For ensemble size M = 2, ridge penalty λ = 0, and any φ ∈ (0,∞),

$$|\mathsf{gcv}_{k,2}^0 - \mathsf{R}_{k,2}^0| \not\xrightarrow{\mathsf{p}} 0.$$

#### Inconsistency on finite ensembles

Proposition 3.3, informal) For ensemble size M = 2, ridge penalty λ = 0, and any φ ∈ (0,∞),

$$|\mathsf{gcv}^{\mathsf{0}}_{k,2} - R^{\mathsf{0}}_{k,2}| \xrightarrow{\mathsf{p}} 0.$$

• The bias scales as 1/M, which is negligible for large *M*:



### Summary

- This work [1] reveals the connections between the *implicit* regularization induced by subsampling and *explicit ridge* regularization for subsample ridge ensembles.
- We establish the uniform consistency of GCV for full ridge ensembles.
- We show that GCV can be *inconsistent* even for ridge ensembles when M = 2.
- Future directions: bias correction of GCV for finite M; extension to other metrics [2]; extension to other base predictors.

[1] Jin-Hong Du, Pratik Patil, and Arun Kumar Kuchibhotla. "Subsample Ridge Ensembles: Equivalences and Generalized Cross-Validation". In: International Conference on Machine Learning (2023)

[2] Pratik Patil and Jin-Hong Du. "Generalized equivalences between subsampling and ridge regularization". In: arXiv preprint arXiv:2305.18496 (2023)